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The Influence of Income Inequality on Air Environmental Quality: An Analysis of the Inland Region in China

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Abstract

The influence of income inequality on air quality is a multifaceted matter that involves society, economy, and the environment. Although the inland region of China has achieved rapid economic development and fully utilized their abundant natural resources, it has also been accompanied by widening income disparities and environmental pollution issues. Especially the frequent occurrence of haze and sandstorms poses serious challenges to the sustainable development and residents' lives in inland areas. To gain a deeper understanding of this issue, this paper is based on panel data from 20 provinces, cities, and autonomous regions in the inland region from 2006 to 2020 and uses the pooled Ordinary Least Squares regression (POLS) to conduct an in-depth analysis of the repercussions of income inequality on air quality. By studying the relationship between the two, this paper aims to provide policy makers with scientific and reasonable policy recommendations, better balance income distribution differences and environmental governance, and attain a sustainable and robust development of the economy and environment in inland areas.

Keywords

Income inequality, Air quality, Pooled ordinary least squares regression, Inland region

Introduction

The inland region of China has become the engine of national development due to the abundant natural resources and key economic status (Zhao & Zou, 2018). The economy has achieved significant development, with a total GDP growth of 837.1% from 6.2 trillion in 2003 to 58.1 trillion in 2022 (the data is from the *Statistical Yearbook of China*, 2000-2020). However, the growth of economy is accompanied by urgent social and environmental challenges, including significant income inequality and air quality issues, which significantly influence residents' quality of life and the sustainable development of society. In addition, the presence of high-income inequality can undermine the strength of institutional frameworks, complicating government efforts to enforce environmental regulations. This challenge arises as wealthier individuals and corporations often possess greater leverage in policy-making processes, potentially leading to a prioritization of their interests over those of the environment. Consequently, governments in such regions may struggle to implement and sustain environmental initiatives, ultimately exacerbating environmental degradation and compromising the effectiveness of their policies (Boyce, 2003).

The harm of income inequality in inland region involves multiple aspects such as economy, society, and environment, which may have serious negative impacts on individuals, society, and countries. Excessive income inequality can easily lead to social conflicts and dissatisfaction, which may lead to social unrest and instability (Hao et al., 2021). Social unrest or political instability, which can, in turn, hinder environmental initiatives (Pata et al., 2022). For instance, social tensions driven by inequality may cause governments to shift their focus and resources away from environmental priorities (Shah et al., 2023). In addition, as the main battlefield of Chinese industrialization, the inland region has injected vitality into economic growth through industrial processes (Zhang et al., 2022). However, this process is accompanied by large-scale industrial emissions, exhaust emissions, and environmental pollution. The decline in air quality has become a prominent issue in inland region, and the frequent occurrence of haze weather has aroused widespread concern (Agarwal et al., 2020). The impact of industrial emissions on the atmosphere, especially the emissions of fine particulate matter and harmful gases, directly threatens the health and living environment of residents (Yousuf et al., 2022). Therefore, studying the influence of income inequality in inland areas on air quality holds profound significance for attaining social equity, environmental justice, and sustainable development goals, which helps to provide more effective policies and action plans for governments, enterprises, and society.

The correlation between income inequality and air quality has always been a concern within the framework of the Sustainable Development Goals (SDGs) (Khan et al., 2022; Zhao et al., 2020). The Sustainable Development Goals provide a clear guiding framework for addressing income inequality and air quality issues in inland areas (Cernev & Fenner, 2020). SDG 10 emphasizes achieving inclusive economic growth at the global, national, and regional levels, reducing inequality among different social groups (Menyelim et al., 2021). Meanwhile, SDG 11 requires the construction of sustainable cities and communities, with a focus on improving the quality of life of residents and protecting the environment (Yamasaki & Yamada, 2022). Placing research within the framework of Sustainable Development Goals (SDGs) helps to better understand and address the interrelationships between society, economy, and environment.

In light of China's long-term goals for 2035, including the improvement of income and wealth distribution patterns alongside the double carbon objectives of reaching a carbon peak by 2030 and achieving carbon neutrality by 2060, addressing the income gap and environmental challenges arising from economic development is of utmost importance (Zhuang & Dou, 2023). The Chinese government faces the urgent task of developing a more equitable and sustainable socioeconomic model. This paper focuses on analyzing how income inequality affects air quality in inland regions. Its objective is to investigate the correlation between narrowing income disparities and enhancing environmental conditions. By doing so, the study aims to offer policymakers evidence-based insights for making informed decisions that align with the Sustainable Development Goals, encompassing economic growth, social equity, and ecological preservation.

Literature Review

Building upon the groundwork laid by Grossman and Krueger (1991), Panayotou (1993) introduced the environmental Kuznets curve (EKC), a pivotal model that analyzes the correlation between income per capita and environmental degradation. This framework has evolved into a foundational cornerstone for exploring the correlation between income and environment issues. Since then, scholars from various backgrounds have delved into this subject, leading to the development of three primary perspectives on the connection between income inequality and air quality.

The negative impact of income inequality on air quality

The prevailing viewpoint, widely embraced by scholars, asserts that income inequality exacerbates air quality issues. Kasuga and Takaya (2017) explored the impact of income inequality on environmental degradation in Japan, establishing not only a correlation but also a causal effect. Their empirical tests revealed a negative impact of income inequality on air quality. Mahalik et al. (2018) delved into the relationship between income inequality and environmental quality across various BRICS economies. Their findings suggested that, apart from South Africa, income inequality contributed to an increase in $CO₂$ emissions. Padhan et al. (2019) observed that the wealth gap decreased $CO₂$ emissions when income inequality exceeded 45%, but expanded $CO₂$ emissions when income inequality was below 40%. They emphasized that countries with more equal income distribution policies experienced a negative impact when exacerbating inequality, whereas countries adopting more unequal income distribution policies to reduce inequality degraded the environment. Baloch et al. (2020) explored the relationship between income inequality and $CO₂$ emissions in 40 Saharan and South African countries from 2010 to 2016. Their empirical results suggested that increasing income inequality exacerbated $CO₂$ emissions, negatively impacting the environmental quality of the study subjects. Alatas and Akin (2022) explored the impact of income inequality on sectoral $CO₂$ emissions in 28 OECD economies, identifying income inequality as a key factor in sectoral emissions. The rise of the Gini index led to varying degrees of increase in $CO₂$ emissions in sectors such as electricity, construction, and transportation. Rai and Rawat (2022) explored the impact of technological innovation, income inequality, and industrialization on the environmental quality of BRICS countries from 1996 to 2016. Their results indicated that a 1% increase in income inequality led to a 1.54% increase in $CO₂$ emissions. Khademolhosseini et al. (2022) investigated the impact of the Gini coefficient on suspended particulate matter and $SO₂$, finding that as the Gini coefficient increased and income inequality intensified, the emissions of suspended particulate matter and $SO₂$ showed varying degrees of increase. Gao and Fan (2023) analyzed the impact of income inequality on CO₂ emissions in the Belt and Road countries, and the results showed that income inequality significantly led to increased carbon emissions and decreased environmental quality.

The positive impact of income inequality on air quality

The alternative viewpoint regarding the correlation posits that an increase in income inequality may contribute to the enhancement of air environmental quality. Grunewald et al. (2017) found that for low- and middle-income economies, higher income inequality was associated with lower carbon emission. Liu et al. (2018) conducted a comprehensive examination of the relationship between income and environmental degradation across 31 provinces in China from 1996 to 2015. According to their research, a certain range of income inequality proved beneficial for enhancing environmental quality. Liu et al. (2019) analyzed the impact of income inequality on carbon emissions in various states of the United States and found that higher income inequality is beneficial for carbon reduction in the long run. Hailemariam et al. (2020) analyzed the impact of income inequality on $CO₂$ emissions in OECD countries and found that an increase in inequality on the Gini index is related to a decrease in CO₂ emissions. Yameogo and Dauda (2022) delved into the intricate relationship between income inequality, carbon emission rates, and economic growth. Their study indicated that income inequality was found to decrease $CO₂$ emissions in Nigeria in the short term. Wan et al. (2022) analyzed data from 217 countries since 1960 and found that rising income inequality can reduce $CO₂$ emissions by reducing energy consumption.

The unclear impact of income inequality on air quality

Some research indicated that the impact of income inequality on environmental outcomes is nuanced, exhibiting variations across different time periods and regions, and in some instances, showing no discernible relationship. Wolde-Rufael and Idowu (2017) found that income inequality was the least important factor determining CO₂ emissions, and there is no direct causal relationship between carbon dioxide emissions and income inequality in these two countries. Hübler (2017) found through fixed effects regression that there was no correlation between income inequality and $CO₂$ emissions. Li (2017) introduced the concept of a threshold effect in the impact of income inequality on the Chinese environment. His theoretical model, supported by empirical evidence, revealed a clear income threshold for China. When a region's per capita income is below the average level, an increase in urban-rural income gap can, to some extent, reduce the emission of environmental pollutants and enhance environmental quality. However, when the per capita income level exceeds the average, it tends to worsen environmental quality. Mader (2018) challenged the conclusions drawn by Knight et al. (2017) and Jorgenson et al. (2017) regarding the promoting effect of income inequality on $CO₂$ emissions. Mader argued that the selected regions, time spans, and indicators to measure inequality in these studies were limited in persuasiveness. In his further research, he found no discernible relationship between social inequality and $CO₂$ emissions. Wang, Yang, et al. (2023) pointed out that under the influence of the threshold effect of income inequality, the relationship between economic growth and per capita carbon emissions presents an N-shaped curve, rather than the traditional inverted U-shaped curve. In the stage of low-income inequality, economic growth has a significant promoting effect on carbon emissions. As income inequality deepens, economic growth in turn suppresses the increase of carbon emissions. However, in the stage of high-income inequality, the impact of economic growth on carbon emissions becomes positive again.

Except direct effects, income inequality may also indirectly affect air pollution through a series of socio-economic mechanisms. These mechanisms include industrial structure (Liu et al., 2020; Wang & Chen, 2024), energy consumption patterns (Duarte et al., 2021; Rojas-Vallejos & Lastuka, 2020; Wang et al., 2021), and environmental regulation (Jorgenson et al., 2017; Qin et al., 2021; Zhou & Li, 2020). Resident health plays an important role as a mediating variable in the process of income inequality affecting air pollution. Air pollution has a serious impact on human health, and the health status of residents is closely related to their income level (Chen & Wang, 2024). Due to limited economic conditions, low-income groups are often more susceptible to the impact of air pollution and have poorer health conditions. And the deterioration of health status will affect their labor ability and income level, further exacerbating income inequality. This vicious cycle makes residents' health an important mediating effect of income inequality on air pollution (Wang, Li, et al., 2023; Yang & Liu, 2018).

Methodology

Selection of EKC econometric model

The convex-concave nature of the EKC shapes the influence of income inequality on environmental quality (Wang $\&$ Lv, 2022). Therefore, it is essential to test the shape of EKC first. The most used econometric models for EKC currently include traditional linear models and logarithmic models. For example:

$$
y_{it} = \alpha_i + \beta_1 \chi_{it} + \beta_2 \chi_{it}^2 + \beta_3 \chi_{it}^3 + \beta_4 z_{it} + \varepsilon_{it}
$$
 (1) (Dinda, 2004)

$$
ln PC_{it} = \alpha_0 + \alpha_1 l n g i n i_{it} + \alpha_2 (l n g i n i_{it})^2 + \alpha_3 (l n g i n i_{it})^3 + \beta l n X_{it} + \mu_i + \nu_t + \varepsilon_{it}
$$

(2) (Zhan, 2018)

This paper uses a linear model to conduct empirical research as follows:

$$
E_{i,t} = \alpha_0 + \beta_1 I N_{i,t} + \beta_2 I N_{i,t}^2 + \beta_3 Z_{i,t} \cdots + \varepsilon_{i,t}
$$
\n(3)

Where, *E* is the dependent variable for measuring environmental quality; *IN* is a variable that measures income inequality; *Z* is the control variable. α_0 is a constant, ε is a random error, where *i* and *t* correspond to province and year.

As shown in Table 1 above, the coefficient sign before the proxy variable of income inequality in Model (3) and the conclusion can provide reference for the judgment of subsequent empirical results.

Coefficient symbol	Conclusion
$\beta_i = \beta_i = 0$	Environmental quality is not affected by income inequality
$\beta_i>0, \beta_i=0$	A positive linear relationship
$\beta_1<0, \beta_2=0$	An inverse linear relationship
$\beta_i>0, \beta_1<0$	An inverted U-shaped curve
$\beta_1<0, \beta_2>0$	A U-shaped curve

Table 1. The Coefficient symbol and conclusion

Source: Author's compilation based on Zaatari (2022)

Data sources and variable descriptions

Data sources

This study utilizes panel data encompassing the Inland region of China (excluding Xizang) over the period from 2006 to 2020. The data sources include the *Statistical Yearbook of China, China Regional Economic Statistical Yearbook, China Energy Statistical Yearbook, China Environmental Statistical Yearbook*, and *Provincial-level Statistical Yearbooks*.

Dependent variable

The primary indicators used for measuring air pollution encompass carbon emissions, $SO₂$ emissions, smoke, and dust emissions, and PM2.5 levels (Luo et al., 2022; Pata, 2021; Zhao et al., 2023). In the contemporary discourse, climate change has gained global attention, with the substantial release of greenhouse gases, predominantly carbon emissions, identified as a leading contributor to this phenomenon. Consequently, many scholarly works have adopted carbon emissions as a key metric for assessing environmental quality (Wolde-Rufael & Idowu, 2017). Presently, China holds the position of the world's largest carbon emitter, placing immense pressure on the nation to reduce emissions. Therefore, this study designates carbon emissions as the principal explanatory variable for evaluating environmental quality.

Unlike indicators such as wastewater, sulphur dioxide, and smoke (powder) dust, there is currently no direct carbon emission statistical data in China. However, the United Nations Intergovernmental Panel on Climate Change (IPCC) provides a standardized method for calculating statistical carbon emissions. The IPCC posits that the surge in carbon emissions is primarily attributable to the utilization of fossil fuels. Hence, this paper adopts the initial method outlined by the IPCC, deemed straightforward and easily calculable. This approach entails estimating CO₂ emissions by considering the quantity of energy fuels consumed along with their respective carbon emission coefficients. Notably, this approach aligns with the common practice among Chinese scholars, as observed in the works of Chen (2009), who predominantly employ this method for estimating carbon emissions across diverse regions in China. For the scope of this research, the same method is applied to estimate carbon emissions data.

The primary energy data required for computing carbon emissions for the 30 provinces from 2006 to 2020 are derived from the *China Energy Statistical Yearbook*. This comprehensive yearbook offers data on eight fossil fuels, encompassing inputs for energy processing and conversion, along with fossil fuels utilized for non-energy purposes. To streamline calculations and avert redundancy, energy balance tables for the 30 provinces from 2006 to 2020 were specifically chosen for computation. This process included excluding inputs and losses occurring during energy processing and conversion procedures, as well as portions used as raw materials and materials in industrial production.

This paper calculates carbon emissions based on provincial terminal energy consumption data over the years. The method is based on the IPCC Guidelines for National Greenhouse Gas Emissions Inventory (IPCC, 2006) . As shown in formula (4):

$$
CE_{it} = \sum (EC_{jit} \times \eta_j)
$$
\n⁽⁴⁾

Where, CE is the total carbon emissions; EC is the *j* energy consumption; η_i is the carbon emission coefficient of the *j* energy source; *i* denotes the province; *t* signifies the year.

The original statistical data on the consumption of diverse energy sources is presented in physical units, necessitating their conversion into standardized statistics when estimating carbon emissions. In accordance with the categorization specified in the China Energy Statistical

Energy Name	Equivalent coal coefficient	Carbon dioxide conversion coefficient
Coal	0.7143kqce/kq	1.9003
Coke	0.9714kqce/kq	2.8604
Crude oil	1.4286kgce/kg	3.0202
Gasoline	1.4714kgce/kg	2.9251
Kerosene	1.4714kgce/kg	3.0179
Diesel oil	1.4571 kgce/kg	3.0959
Fuel oil	1.4286kgce/kg	3.175
Liquefied petroleum gas	1.7143kgce/kg	3.1013
Natural gas	1.33 kgce/m ³	21.622

Table 2. The equivalent coal coefficient and conversion coefficient of 9 types of energy

Source: General Principles for Calculation of Comprehensive Energy Consumption (GB/T 2589-2020), 2021

Yearbook, the final energy consumption is divided into nine categories: coal, coke, crude oil, gasoline, kerosene, diesel, fuel oil, liquefied petroleum gas, and natural gas. The corresponding equivalent coal coefficients and conversion coefficients for these nine energy types are detailed in Table 2 below.

In addition, SO_2 is one of the main atmospheric pollutants formed by industrial fuel combustion, and fuel use is usually considered as the environmental effect of energy consumption. Therefore, the robustness testing section of this paper will replace carbon emissions with sulfur dioxide indicators.

Independent Variables

At present, the main indicators for measuring income inequality include the Gini coefficient, the Thiel index, and the urban-rural income ratio (Consoli et al., 2023; Tang et al., 2022). This paper uses the Thiel index to measure income inequality.

The value of the Thiel index is positively correlated with income inequality, that is, the larger the Thiel index, the greater the income inequality. This paper refers to the specific calculation formula used in the Wang and Ouyang's research for the calculation method of the Thiel index (Wang & Ouyang, 2008).

The formula is as follows:

$$
Theil_{it} = \sum_{j=1}^{2} \left(\frac{I_{it,j}}{I_{it}} \right) ln \left(\frac{I_{it,j}}{P_{it,j}} / \frac{I_{it}}{P_{it}} \right)
$$
\n
$$
\tag{5}
$$

Where, *j* takes values 1 and 2 representing urban and rural areas, respectively; *Theil* is the Theil index; *I* represents the total income, calculated by multiplying the total population by per capita income; *P* is the total population; *i* denotes the province; t signifies the year.

The Theil index, computed with this formula, incorporates the population shares of urban and rural areas as weights, considering the evolving urban and rural population structures. This approach is particularly apt for capturing the nuances of the current income landscape amid the dual structure of urban and rural areas in China. One of the salient features of the Theil index lies in its versatility to encapsulate both "inter-regional" and "intra-regional" income disparities. At the inter-regional level, it adeptly measures the variations in average income levels across diverse geographical areas. Conversely, delving into the regional level, the Theil index dissects income gaps among individuals or subgroups within a particular region. This dual-faceted decomposition capability renders the Theil index an invaluable tool for exploring multi-layered and multidimensional income inequality dynamics (Heng, 2011).

This paper primarily delves into the influence of inter-regional income distribution on air quality. One plausible reason for this focus stems from the practical challenges associated with gathering intra-regional income distribution data, particularly when encompassing vast numbers of individuals or segmented groups. To ensure the study's feasibility and efficiency, the article likely opted for the more readily accessible and manageable inter-regional data.

Moreover, the study's emphasis rests on elucidating the macro-level impact of income inequality on air quality. Consequently, regional disparities, as a pivotal aspect of macroinequality, naturally emerged as the central research focus. Additionally, regional disparities encompass broader policy, economic, and social dimensions, which are crucial for comprehending the intricate relationship between overall inequality and air quality.

Nonetheless, this methodological choice bears notable limitations. Firstly, overlooking intraregional income distribution may skew the overall inequality picture, potentially under- or overestimating it, thereby compromising the precision of research findings. Secondly, intraregional income inequality might exert its influence on air quality through distinct mechanisms, such as shaping consumption habits or influencing environmental policy implementation, which might remain elusive in an inter-regional analysis.

Future research should aim to include more detailed geographical data and develop multi-level Theil indices. This would enable a comprehensive view of income distribution's influence on air quality, uncovering cross-strata relationships and informing targeted policies.

Control variables

This paper mainly uses economic development level, urbanization level, industrial structure indicators, and export dependence as control variables.

Economic development level. Economic growth, reflected in increased residents' disposable income, exerts influence on environmental quality by impacting GDP and environmental investment. In this study, the economic development level of each region is represented by per capita GDP. The issues of income disparity and environmental quality are inherently linked to economic development, with a discernible relationship existing among the level of economic development, income inequality, and environmental outcomes. The necessary GDP and population data are sourced meticulously from the *Statistical Yearbook of China*.

Urbanization level. The degree of urbanization plays a crucial role in influencing environmental quality. In contrast to developed nations, China is presently undergoing a phase of swift urbanization, marked by a substantial migration of rural residents to urban areas. However, China's Hukou system has led to the underestimate of the population actually living in the city but not moving in with registered residence, which affects the accuracy of urbanization data (Wang et al., 2022). This deviation in data further affects the accurate assessment of environmental effects such as environmental pressure and resource consumption caused by urbanization. In future research, multi-channel data collection and comprehensive analysis should be conducted to more comprehensively evaluate the level of urbanization and its impact on environmental quality. Some scholars believe that the relationship between urbanization and environmental pollution may not be single linear, but rather that with the development of urbanization, environmental quality will undergo a process of first deteriorating and then improving, which will present an inverted U-shape. This paper gauges the degree of urbanization development by calculating the percentage of the urban population in various regions, expressed as a proportion of the total population at the year's conclusion. The essential data for this measurement is diligently sourced from the *Statistical Yearbook of China*.

Industrial structure indicator. Industrial structure factors constitute important variables, given their close connection to energy consumption, which, in turn, influences environmental quality (Li & Yang, 2017). Regions with a higher proportion of the industrial sector tend to exhibit greater energy consumption, resulting in more pronounced environmental pollution. Prevailing research on industrial structure and environmental quality indicates that an excessive share of the secondary industry contributes to an imbalanced industrial structure, exacerbating environmental issues. Analyzing the repercussions of industrial upgrading on the ecological environment within the Yangtze River Economic Belt, Li (2018) contends that the rationalization and elevation of industrial structure have a positive impact on enhancing the ecological environment. This underscores a symbiotic relationship between industrial upgrading and environmental amelioration. This paper utilizes the percentage of added value from the secondary industry in the province's GDP (*psi*) as the indicator for assessing industrial structure. Data for this analysis is derived from the National Bureau of Statistics website and the *Statistical Yearbook of China*.

Export dependency. There is still no consensus on the impact of export dependence on environmental quality. Some scholars believe that with the increasing openness of export trade, enterprises continuously expand their production scale to seize the international market, resulting in environmental damage (Xue & Ding, 2020). However, there are also studies that hold different perspectives. Some scholars believe that in recent years, international trade competition has intensified, and countries have raised the admission standards for imported goods and formulated strict environmental regulations, especially developed countries that have higher requirements for environmental protection. Therefore, to avoid export risks, Chinese export enterprises are poised to advocate for clean production and prioritize the development of environmentally friendly products through innovations in environmental technology. On the one hand, this can ensure that the product meets the environmental access standards established by the exporting country, reduce restrictions on China's export goods, increase export volume to compensate for increased environmental costs, and on the other hand, it is conducive to enhancing the environmental awareness of enterprises in China and improving the environmental quality of China (Chen, 2012). In addition, foreign direct investment also affects environmental conditions through various factors, such as the introduction of new technologies and environmental concepts, which directly affect the level of environmental pollution control and thus the degree of environmental pollution. Li et al. (2017) believed that foreign direct investment has promoted the development of China's scientific and technological level, especially environmental protection technology, by exerting its "demonstration effect" and "spillover effect" on pollution control and has a restraining effect on China's environmental pollution. However, according to Han et al. (2019), while foreign direct investment (FDI) may stimulate short-term enhancements in China's environmental quality, its long-term effects are characterized by a suppressive influence on sustained environmental improvement, ultimately yielding a negative impact on overall environmental quality. Therefore, this paper uses the proportion of total import and export volume of foreign-invested enterprises to GDP to measure export dependence. The total import and export figures for foreign-invested enterprises are adjusted using the average exchange rate for the respective year. All data originate from the *Statistical Yearbook of China*.

Detailed descriptions of the variables are presented in Table 3 below.

Descriptive statistics

To have a preliminary understanding of the basic characteristics of the variables used in the model, Table 4 below provides the descriptive statistical characteristics of each variable.

Table 3. Description of related variables

Source: Author's compilation

Table 4. Statistical description of variables

Source: Stata Corp2021

Findings

Empirical findings

This section first uses the pooled Ordinary Least Squares regression (pooled OLS) to empirically test the impact in the inland region, and then conducts a robustness testing on the measurement results.

Regression results

As shown in Table 5 above, the coefficient of the *thirl* is 218.525, which is a positive number, while the coefficient of its square term (*thirl²*) is -395.191, reflecting a negative value. Both coefficients have passed the 1% significance test. This indicates that in inland region, there is an inverted U-shaped relationship between *pce* and the *thirl*, which means there is an inverted U-shaped EKC curve in inland China.

Furthermore, further observation of the regression results of the control variables revealed that:

Per capita GDP (*pgdp*): The results of the regression analysis reveal an inverse relationship between *pgdp* and environmental pollution in inland region, that is, as *pgdp* increases, per capita carbon emissions (*pce*) will decrease, and environmental quality will improve.

Urbanization rate (*ur*): The outcomes of the regression analysis indicate a negative impact of the urbanization rate on environmental quality in inland areas, that is, the higher the level of *ur*, the more pollutants are emitted.

Per capita added value of the secondary industry (*psi*): The regression analysis results show a positive relationship between *psi* and environmental pollution in inland areas, that is, as *psi* increases, the amount of pollutant emissions increases.

The proportion of total import and export volume of foreign-invested enterprises to GDP (*tied*): The regression analysis results show a positive relationship between *tied* and environmental quality in inland areas, that is, as *tied* increase, environmental quality gradually improves.

Robustness testing

To enhance the credibility of the research findings, this paper performs robustness testing using four distinct methods. The first method is replacing the dependent variable, the second method is adding variable, the third method is reducing variable, and the fourth method uses a lag of one period for robustness detection.

Replacing variables

As shown in Table 6, after replacing the dependent variable, the regression results and significance of the independent variables align closely with the regression results (Table 5). However, there are slight variations in the magnitudes of the coefficients and the extent of their

Table 5. Regression results

****p*<0.01, ***p*<0.05, * *p*<0.1

Source: Stata Corp2021

pso	Coef.	St.Err.	t-value	p-value	[95% Conf Interval]	Sig
thirl	0.396	0.072	5.51	0	0.254	$* * *$ 0.537
thirl2	-0.732	0.215	-3.40	0.001	-1.156	-0.309
pgdp	-0.014	0.002	-9.07	0	-0.017	*** -0.011
ur	0.15	0.016	9.49	$\mathbf{0}$	0.119	*** 0.181
psi	0.02	0.003	5.90	0	0.014	0.027
tied	-0.077	0.017	-4.51	0	-0.11	-0.043
Constant	-0.073	0.011	-6.81	$\mathbf{0}$	-0.095	-0.052

Table 6. The result with pso as the dependent variables

****p*<0.01, ***p*<0.05, * *p*<0.1

Source: Stata Corp2021

Table 7. Increasing ratio of completed investment in waste gas treatment projects (rwg)

pce	Coef.	St.Err.	t-value	p-value	[95% Conf Interval]	Sig
thirl	177.037	40.911	4.33	$\mathbf 0$	96.519	257.554
thirl2	-335.789	121.731	-2.76	0.006	-575.371	-96.207 ***
pgdp	-2.871	0.924	-3.11	0.002	-4.689	-1.052
ur	92.027	9.343	9.85	$\mathbf 0$	73.638	110.416 ***
psi	7.375	2.019	3.65	$\mathbf 0$	3.401	11.349 ***
tied	-63.504	10.03	-6.33	$\mathbf 0$	-83.244	-43.765 ***
rwg	2867.209	355.924	8.06	Ω	2166.708	3567.71
Constant	-53.448	6.203	-8.62	$\mathbf 0$	-65.657	-41.239

****p*<0.01, ***p*<0.05, * *p*<0.1

Source: Stata Corp2021

impact. The coefficient of *thirl* is 0.396, which is positive, while the coefficient of its square term is -0.732, which is negative, and both pass the 1% significance test.

Adding variables

As shown in Table 7, after adding the control variable *rwg*, all variables passed the 1% significance test.

Reducing variables

As shown in Table 8, after reducing the control variable *tied*, *pgdp and psi* exhibited significance at the 5% level, while other variables demonstrated significance at the 1% level.

Lagged regression results for one period

As shown in Table 9, after a lag of one period, all variables passed the 1% significance test.

The robustness testing conducted through the four methods affirms the overall reliability of the empirical results presented in this paper.

pce	Coef.	St.Err.	t-value	p-value	[95% Conf Interval]	Sig
thirl	285.529	48.822	5.85	0	189.443	381.615
thirl2	-566.313	146.842	-3.86	0	-855.308	-277.317 ***
pgdp	-2.184	.99	-2.21	0.028	-4.132	-0.236 **
ur	94.942	10.56	8.99	0	74.159	115.724 ***
psi	5.674	2.264	2.51	0.013	1.218	10.129 **
Constant	-63.919	7.445	-8.59	0	-78.572	-49.266

Table 8. Reducing the proportion of total import and export volume of foreign-invested enterprises to GDP (*tied*)

****p*<0.01, ***p*<0.05, * *p*<0.1

Source: Stata Corp2021

pce	Coef.	St.Err.	t-value	p-value	[95% Conf Interval]	Sig
thirl	206.271	43.78	4.71	Ω	120.074	*** 292.468
thirl2	-369.236	135.413	-2.73	0.007	-635.85	-102.622
pgdp	-5.233	0.9	-5.82	Ω	-7.004	-3.462
ur	112.9	9.414	11.99	Ω	94.365	131.436 ***
psi	11.597	1.996	5.81	Ω	7.667	15.527 ***
tied	-88.619	10.113	-8.76	0	-108.53	-68.709 [*]
Constant	-61.951	6.425	-9.64	Ω	-74.601	*** -49.301

Table 9. Lagged regression results for one period

****p*<0.01, ***p*<0.05, * *p*<0.1

Source: Stata Corp2021

Regression fitting graphs

As shown in Figure 1 above, during the quadratic fitting of *pce* and *thirl*, the estimated coefficient of *thirl* is 58.333, while the estimated coefficient of *thirl²* is -308.66. The fitting results indicate the existence of an inverted U-shaped EKC curve in inland China.

As shown in Figure 2 below, the quadratic fitting results of *pce* and *pgdp* indicate an inverted U-shaped curve between *pce* and *pgdp* in inland areas. Based on the quadratic fitting function of *pce* and *pgdp*, *pce* = 1.4797 + 3.5973 *pgdp* - 0.23466 *pgdp²* . Setting *pce'* = 0, i.e., 3.5973 - 0.46932 *pgdp* = 0, obtain: *pgdp* = 7.6649, which means that the inflection point of *pce* is around *pgdp* equals 7.6649. The *pgdp* in 2018 was 5.0895, 5.4697 in 2019, and 5.6304 in 2020 (data from 2018 to 2020 is calculated based on *the Statistical Yearbook of China*). This indicates that *pgdp* is in the ascending stage to the left of the inflection point, that is, as *pgdp* increases, *pce* will increase.

As shown in Figure 3 below, the quadratic fitting results of *pce* and *ur* indicate a U-shaped relationship between *pce* and *ur* in inland areas. Following the approach of calculating the inflection point above, the inflection point of *ur* is 29.68%. In 2018, the *ur* was 56.59%, 57.98% in 2019, and 59.23% in 2020 (data from 2018 to 2020 is calculated based on *the Statistical Yearbook of China*). It is evident that the *ur* has surpassed the inflection point and is currently in an ascending phase on its right, that is, as *ur* increases, *pce* will increase.

Figure 1. Quadratic fitting graph between *pce* and *thirl* in inland region Source: Stata Corp2021

Figure 2. Quadratic fitting graph between *pce* and *pgdp* Source: Stata Corp2021

As shown in Figure 4 above, the quadratic fitting results show an inverted U-shaped relationship between *pce* and *psi* in inland areas. Following the approach of calculating the

Figure 3. Quadratic fitting graph between *pce* and *ur* Source: Stata Corp2021

Figure 4. Quadratic fitting graph between *pce* and *psi* Source: Stata Corp2021

inflection point, the inflection point of *psi* is 2.7226. The *psi* in 2018 was 2.0555, 2.182 in 2019, and 2.1715 in 2020 (the data from 2018 to 2020 is calculated based on *the Statistical Yearbook of*

Figure 5. Quadratic fitting graph between *pce* and *tied* Source: Stata Corp2021

China). It is obvious that *psi* has not yet reached the inflection point and is in the rising stage on its left side, that is, as *psi* increases, *pce* will increase.

As shown in Figure 5 below, the quadratic fitting results of *pce* and *tied* indicate a U-shaped relationship between *pce* and *tied* in inland areas. Following the approach of finding the inflection point, it can be concluded that the inflection point is 10.23%. In 2018, the *tied* was 5.21%, 4.88% in 2019, and 5.09% in 2020 (data from 2018 to 2020 is calculated based on *the Statistical Yearbook of China*). It is obvious that the *tied* has not yet reached the inflection point, which is the descending stage on its left side, that is, as the *tied* increases, the *pce* will decrease.

Conclusions and Implications

In general, an inverted U-shaped EKC exists in inland China, which first increases and then decreases. This indicates that the EKC hypothesis is applicable in inland China.

Conclusions

The regression analysis and quadratic fitting results consistently indicate an inverted U-shaped relationship between the environmental quality and the income inequality.

Per capita GDP: The regression analysis results reveal a negative correlation between per capita GDP and air environmental pollution in inland areas, that is, as per capita GDP increases, per capita carbon emissions will decrease, and air environmental quality will improve. Furthermore, by further calculating the inflection point, it was found that per capita GDP is in the rising stage to the left of the inflection point of the fitting curve.

Urbanization rate: The regression analysis results show that urbanization rate has a negative

effect on environmental quality in inland areas, that is, the higher the level of urbanization rate, the more pollutants are emitted. The quadratic fitting graph illustrates a U-shaped relationship between urbanization rate and air environmental quality.

Per capita added value of the secondary industry: The regression analysis results show a negative relationship between per capita added value of the secondary industry and air environmental quality in inland areas. The quadratic fitting graph illustrates an inverted U-shaped curve.

The proportion of total import and export volume of foreign-invested enterprises to GDP: The regression analysis results show a positive relationship between *tied* and environmental quality in inland areas, that is, as the proportion of total import and export volume of foreign-invested enterprises to GDP increase, environmental quality gradually improves.

Implications

Strengthening green technology to promote clean energy development

Enhancing green technology and implementing clean energy policies represent crucial strategies in addressing both income inequality and environmental quality concerns (Liu et al., 2022). It is imperative for the government to actively promote, and support research and application of green technologies aimed at reducing environmental pollution and resource consumption. Investment in clean energy sources, such as solar and wind energy, is essential for mitigating reliance on highly polluting energy. Moreover, increased funding in green technologies and clean energy infrastructure, encompassing power networks, energy storage facilities, and charging stations, can significantly contribute to the sustainability of clean energy (Razmjoo et al., 2021). Furthermore, boosting investment in research and development within the realms of green technology and clean energy is vital for expediting the development and commercialization of new technologies (Kong, 2022). This can be achieved through government funding, collaborative research and development initiatives, and innovation funds.

Enhancing education and skill training

Reinforcing education and skill training policies stands as a crucial strategy for mitigating income disparities and enhancing environmental quality in inland areas. Elevating the education level and skill literacy of residents not only fosters increased employment opportunities but also facilitates the creation of high-value industries. Moreover, it cultivates a deeper understanding and a heightened sense of responsibility toward the environment among individuals (Wei et al., 2020).

Facilitating local government cooperation

Facilitating collaboration among governments at various levels is pivotal for coordinating the execution of environmental policies, especially concerning cross-regional environmental issues such as watershed management and air quality improvement. The establishment of cross-regional planning agencies can effectively coordinate development plans and policies across different regions, ensuring the judicious allocation of resources and mitigating inequality (Gao et al., 2022). Encouraging local governments to share resources and best practices promotes sustainable development and environmental protection initiatives. Collaborative efforts on cross-regional environmental protection projects, encompassing water resource management, air pollution control, and ecological preservation, can collectively address environmental challenges and enhance environmental quality (He et al., 2022).

Setting long-term plans and goals

To attain the objective of diminishing income inequality and enhancing environmental quality over the long term, the government can craft comprehensive, enduring plans and policies with a focus on nurturing the tourism industry in inland areas (Xiao et al., 2022). Inland regions, characterized by abundant natural and cultural tourism resources, have experienced relatively sluggish development in this sector due to factors such as lagging overall development and inconvenient transportation (Ma et al., 2023). To alleviate income inequality, stimulate economic growth in inland areas, and enhance environmental quality, the government can implement a series of policies aimed at fostering the development of the inland tourism industry.

In conclusion, to narrow the income disparity in inland regions while concurrently enhancing environmental quality, the government should enact comprehensive policies. In inland areas, emphasis can be placed on bolstering infrastructure construction, fostering enterprise investment, advancing tourism development, fortifying education and training, and elevating residents' skill levels. Simultaneously, there should be a concerted effort to prioritize environmental protection and formulate sustainable development plans.

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